

Markov Chain Monte Carlo Methods: Gibbs Sampling

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Basic MCMC: Gibbs sampling for bivariate PDF

Gibbs sampling for a bivariate PDF $f_{XY}(x, y)$

Assume that the two conditional PDFs $f_{Y|X}(y|x)$ and $f_{X|Y}(x|y)$ are (easily) sampleable.

Start with an arbitrary initial value x_1 .

① Sample $y_1 \sim f_{Y|X}(\cdot|x_1) \implies (x_1, y_1)$

② Sample $x_2 \sim f_{X|Y}(\cdot|y_1)$

③ Sample $y_2 \sim f_{Y|X}(\cdot|x_2) \implies (x_2, y_2)$

④ Sample $x_3 \sim f_{X|Y}(\cdot|y_2)$

⑤ Sample $y_3 \sim f_{Y|X}(\cdot|x_3) \implies (x_3, y_3)$

⑥ ...

Repeat as required.

Basic MCMC: Gibbs sampling for bivariate PDF

Expressing the previous algorithm concisely

- 1 Start with an arbitrary initial value x_1 .
- 2 (x_{i+1}, y_{i+1}) is obtained from (x_i, y_i) by sampling from the conditionals:

$$\begin{aligned}y_{i+1} &\sim f_{Y|X}(\cdot|x_i) \\x_{i+1} &\sim f_{X|Y}(\cdot|y_{i+1})\end{aligned}$$

- 3 Repeat step 2 as required.

Example: Gibbs sampling for a bivariate normal density

Let

$$\mathbf{x} = \begin{pmatrix} x \\ y \end{pmatrix} \quad \mu = \begin{pmatrix} \mu_X \\ \mu_Y \end{pmatrix} \quad \Sigma = \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}.$$

Let us consider the bivariate normal PDF

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{2\pi} |\Sigma|^{-1/2} \exp \left[-\frac{1}{2} (\mathbf{x} - \mu)^T \Sigma^{-1} (\mathbf{x} - \mu) \right]$$

where Σ^{-1} is the matrix inverse of Σ :

$$\Sigma^{-1} = \frac{1}{1 - \rho^2} \begin{bmatrix} 1 & -\rho \\ -\rho & 1 \end{bmatrix}.$$

The determinant $|\Sigma| = 1 - \rho^2$.

Example: Gibbs sampling for a bivariate normal density

Marginal PDFs

$$f_X(x) \equiv N(\mu_X, 1)$$

$$f_Y(y) \equiv N(\mu_Y, 1)$$

Example: Gibbs sampling for a bivariate normal density

Conditional PDFs

$$f_{X|Y}(x|y) = \frac{1}{\sqrt{2\pi(1-\rho^2)}} \exp\left[-\frac{x-\rho y}{2(1-\rho^2)}\right] \equiv N(\rho y, 1-\rho^2)$$
$$f_{Y|X}(y|x) = \frac{1}{\sqrt{2\pi(1-\rho^2)}} \exp\left[-\frac{y-\rho x}{2(1-\rho^2)}\right] \equiv N(\rho x, 1-\rho^2)$$

We know how to sample from these univariate normal PDFs ...

<https://www2.stat.duke.edu/courses/Spring12/sta104.1/Lectures/Lec22.pdf>

Example: Gibbs sampling for bivariate normal

```
rbvnorm.gibbs <- function( n, x1, mu = c( 0, 0 ), rho = 0 )
{
  # Gibbs sampler for a correlated bivariate normal N( mu, Sigma2 )
  # where
  # mu = [ mu1 mu2 ]
  # and
  # covariance matrix Sigma^2 = [ 1  rho ]
  #                               [ rho  1 ]

  stopifnot( abs( rho ) < 1, length( mu ) == 2, is.numeric( mu ) )

  r <- matrix( nrow = n, ncol = 2 )

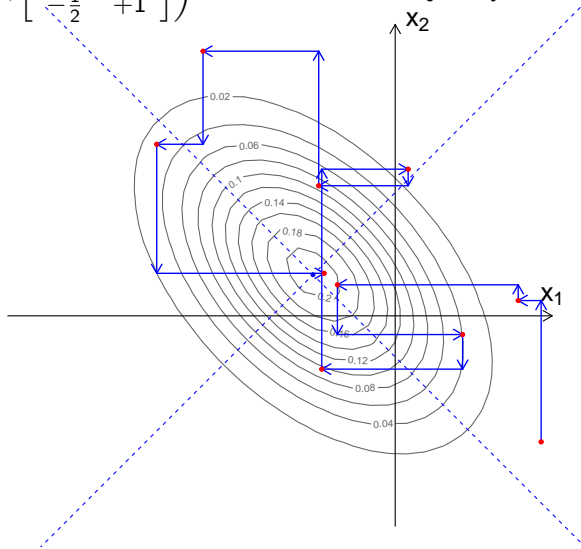
  sd <- sqrt( 1 - rho^2 )

  r[1,1] <- x1
  r[1,2] <- rnorm( 1, mean = rho * r[1,1], sd = sd )
  for ( i in 2:n )
  {
    r[i,1] <- rnorm( 1, mean = rho * r[i-1,2], sd = sd )
    r[i,2] <- rnorm( 1, mean = rho * r[i, 1], sd = sd )
  }

  cbind( r[,1] + mu[1], r[,2] + mu[2] )
}
```

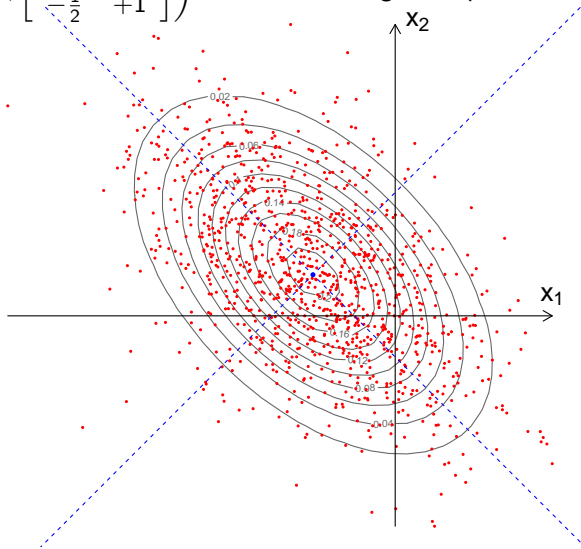
Example: Gibbs sampling for bivariate normal

$N\left(\left(-1, +\frac{1}{2}\right), \begin{bmatrix} +1 & -\frac{1}{2} \\ -\frac{1}{2} & +1 \end{bmatrix}\right) :: \text{Gibbs} :: \text{Initial trajectory}$



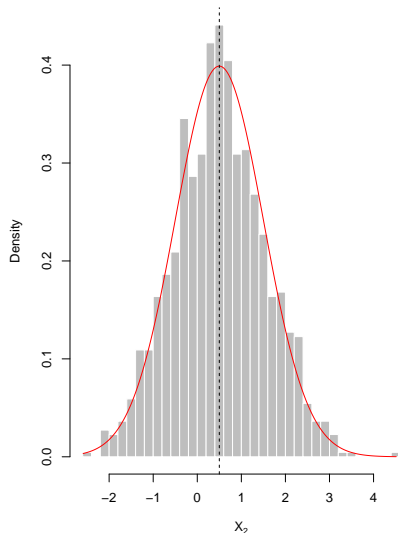
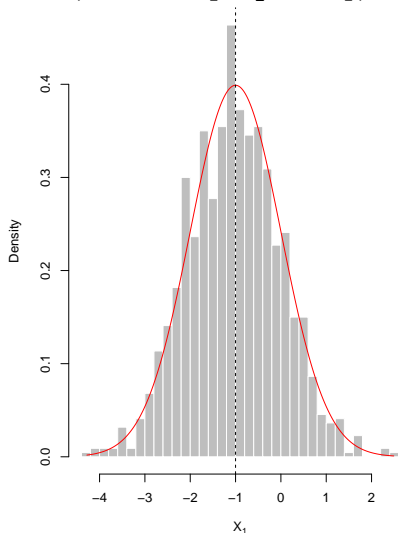
Example: Gibbs sampling for bivariate normal

$N\left(\left(-1, +\frac{1}{2}\right), \begin{bmatrix} +1 & -\frac{1}{2} \\ -\frac{1}{2} & +1 \end{bmatrix}\right) :: \text{Gibbs} :: \text{A larger sample}$



Example: Gibbs sampling for bivariate normal

$N\left(\left(-1, +\frac{1}{2}\right), \begin{bmatrix} +1 & -\frac{1}{2} \\ -\frac{1}{2} & +1 \end{bmatrix}\right) :: \text{Gibbs} :: \text{Marginal histograms}$



Example: Gibbs sampling for bivariate normal

Suppose that the Gibbs sampling trajectory is

$$x_1 \rightarrow y_1 \rightarrow x_2 \rightarrow y_2 \rightarrow \dots \rightarrow x_t \rightarrow y_t \rightarrow \dots$$

- Time evolution of the joint PDF of (X, Y) under this Markov chain:

$$N \left(\begin{bmatrix} \mu_X + \rho^{2t} x_1 \\ \mu_Y + \rho^{2t+1} x_1 \end{bmatrix}, \begin{bmatrix} 1 - \rho^{4t} & \rho(1 - \rho^{4t}) \\ \rho(1 - \rho^{4t}) & 1 - \rho^{4t} \end{bmatrix} \right)$$

for $|\rho| < 1$.

- As $t \rightarrow \infty$, this becomes

$$N \left(\begin{bmatrix} \mu_X \\ \mu_Y \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} \right)$$

which is the target PDF.

in the handout
on the web

...

Gibbs sampling for $f_{XYZ}(x, y, z)$. Assume that the conditionals

$$f_{Z|XY}(z|x, y), f_{Y|XZ}(y|x, z), f_{X|YZ}(x|y, z)$$

are (easily) sampleable.

- 1 Start with arbitrary initial values x_1, y_1 .
- 2 $(x_{i+1}, y_{i+1}, z_{i+1})$ is obtained from (x_i, y_i, z_i) by sampling from the conditionals:

$$z_{i+1} \sim f_{Z|XY}(\cdot|x_i, y_i)$$

$$y_{i+1} \sim f_{Y|XZ}(\cdot|x_i, z_{i+1})$$

$$x_{i+1} \sim f_{X|YZ}(\cdot|y_{i+1}, z_{i+1})$$

- 3 Repeat step 2 as required.

Generalizing Gibbs for a multivariate PDF

Suppose

$f(x_1, \dots, x_k)$: Target multivariate PDF

$f_i(x_i | x_{-i})$: conditional PDF of x_i given everything else (x_{-i})

New notation $x_{-i} \equiv$ “all variables except the i th”

Gibbs for a multivariate PDF: Sequential scan

- 1 Start with arbitrary initial values x_1, \dots, x_k .
- 2 For $i = 1, \dots, k$:
Sample $x_i \sim f_i(\cdot | x_{-i})$
- 3 Repeat step 2 as required.

Gibbs for a multivariate PDF: Random scan

- 1 Start with arbitrary initial values x_1, \dots, x_k .
- 2 For $i \in$ a random permutation of $1, \dots, k$:
Sample $x_i \sim f_i(\cdot | x_{-i})$
- 3 Repeat step 2 as required.

- The basic variant discussed changes one variable at a time.
- Block versions, where the conditionals are expressed over groups of variables instead of single variables, are also possible.

- Gibbs algorithm sets up a Markov chain, because the next state depends only on the current state.
- Gibbs can be shown to be a special case of Metropolis-Hastings with acceptance probability = 1.

Gibbs as Metropolis-Hastings with $P(\text{accept})=1$

- $f(x_1, \dots, x_k)$: Target multivariate PDF
- $f_i(x_i|x_{-i})$: conditional PDF of x_i given everything else :: x_{-i}
- Suppose the current state of the Markov chain is (x_1, \dots, x_k) .
- The Gibbs sampler is now ready to change the i th variable x_i .
- Think of Gibbs as Metropolis-Hastings with proposal PDF

$$q(\cdot|x_i, x_{-i}) = f_i(\cdot|x_{-i}).$$

- Generate a candidate from this proposal PDF,

$$x'_i \sim f_i(\cdot|x_{-i}).$$

Gibbs as Metropolis-Hastings with $P(\text{accept})=1$

Recall the Metropolis-Hastings form of the acceptance probability

$$a(\text{proposed}|\text{current}) = \min \left\{ 1, \frac{f(\text{proposed})}{f(\text{current})} \times \frac{q(\text{current}|\text{proposed})}{q(\text{proposed}|\text{current})} \right\}$$

In our Gibbs sampling setting,

- Current state of the chain: (x_i, x_{-i})
- Proposed state of the chain: (x'_i, x_{-i})

Gibbs as Metropolis-Hastings with $P(\text{accept})=1$

For convenience, let us write $f(x_1, \dots, x_k)$ as $f(x_i, x_{-i})$.

M-H acceptance probability for Gibbs sampling is

$a(x'_i, x_{-i} | x_i, x_{-i}) = \min(1, \alpha)$, where

$$\begin{aligned}\alpha &= \frac{f(x'_i, x_{-i})}{f(x_i, x_{-i})} \times \frac{q(x_i, x_{-i} | x'_i, x_{-i})}{q(x'_i, x_{-i} | x_i, x_{-i})} \\ &= \frac{f(x'_i, x_{-i})}{f(x_i, x_{-i})} \times \frac{f_i(x_i | x_{-i})}{f_i(x'_i | x_{-i})} \\ &= \frac{f_i(x'_i | x_{-i}) f(x_{-i})}{f_i(x_i | x_{-i}) f(x_{-i})} \times \frac{f_i(x_i | x_{-i})}{f_i(x'_i | x_{-i})} \\ &= 1.\end{aligned}$$

That is, $a(x'_i, x_{-i} | x_i, x_{-i}) = 1$. This is true for any i .

Gibbs as Metropolis-Hastings with $P(\text{accept})=1$

Convergence, caveats, and burn-in/equilibration considerations applicable to M-H are also applicable to Gibbs – except for tunable parameters in the proposal PDF of M-H.

- A classic
Explaining the Gibbs Sampler. Casella & George, The American Statistician, Vol. 46, No. 3 (Aug., 1992), pp. 167-174. <http://www.jstor.org/stable/2685208>
- A perspective on comparison between Gibbs and M-H:
<https://stats.stackexchange.com/questions/104815/gibbs-sampling-versus-general-mh-mcmc>
- Advanced texts with rigorous treatment of the subject
 - Monte Carlo Statistical Methods. Robert & Casella, Springer 2004
 - Advanced Markov Chain Monte Carlo Methods. Liang, Liu, & Carroll, Wiley 2010